Incidence of Diabetes Following Ramipril or Rosiglitazone Withdrawal

THE DREAM TRIAL INVESTIGATORS*

OBJECTIVE—To examine the impact of withdrawing rosiglitazone and ramipril medication on diabetes incidence after closeout of the Diabetes REduction Assessment with ramipril and rosiglitazone Medication (DREAM) trial.

RESEARCH DESIGN AND METHODS—The 3,366 DREAM subjects at trial end who had not developed diabetes while taking double-blind study medication were transferred to single-blind placebo for 2 to 3 months before undergoing an oral glucose tolerance test. Glycemic status was analyzed for the trial plus washout period and for the washout period alone.

RESULTS—Following median (interquartile range) 71 (63–86) days drug withdrawal, overall glycemic status remained modestly improved in those allocated ramipril during the trial with an 11% increase in regression to normoglycemia, compared with placebo. In those previously allocated rosiglitazone, glycemic status remained substantially improved with a 49% reduction of new-onset diabetes or death and a 22% increase in regression to normoglycemia, compared with placebo. However, during the washout phase alone the incidence of diabetes or death was identical for those allocated previously to ramipril or placebo, or to rosiglitazone or placebo.

CONCLUSIONS—In people allocated to ramipril compared with those not allocated ramipril during the trial, the postwashout normoglycemia incidence was higher. In people allocated to rosiglitazone compared with those not allocated rosiglitazone during the trial, the postwashout incidence of diabetes was significantly lower and the incidence of normoglycemia was higher. During the washout period, diabetes incidence was the same for ramipril versus placebo and for rosiglitazone versus placebo. Rosiglitazone delays disease progression during treatment but the process resumes at the placebo rate when the drug is stopped.

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ype 2 diabetes prevention is a major public health issue. The Diabetes REduction Assessment with ramipril and rosiglitazone Medication (DREAM) trial (1) showed that the new-onset diabetes rate in people with impaired glucose tolerance (IGT) and/or impaired fasting glucose (IFG) can be slowed substantially (60% relative to placebo) by the thiazolidinedione rosiglitazone (2) but not by the ACE inhibitor ramipril (3). The rosiglitazone results were similar to those seen in two trials with the now withdrawn thiazolidinedione troglitazone (4,5). DREAM also showed that both rosiglitazone and ramipril increased the rate of regression to normal glucose tolerance

Because type 2 diabetes is a progressive disorder (6), it is of interest to know

whether an intervention that delays diabetes onset is 1) masking diabetes appearance by suppressing glucose levels, 2) slowing diabetes development only while it is being administered, or 3) has sustained benefit even after withdrawal. These possibilities can be distinguished by washing out the intervention and then reassessing glycemic status (Fig. 1). If diabetes was being masked, postwashout diabetes incidence in the intervention group would exceed that in the placebo group and the overall (trial plus washout) diabetes prevalence in both groups would be similar (Fig. 1, scenario 1). If the underlying disease process only slowed during intervention, postwashout diabetes incidence would be similar in the intervention and placebo groups, with the overall prevalence of diabetes in the group

formerly receiving the intervention remaining lower than in the placebo group (Fig. 1, scenario 2). Finally, if the intervention had a sustained effect after drug withdrawal, both the washout incidence and the overall diabetes prevalence in the group formerly receiving the intervention would be lower than in the placebo group (Fig. 1, scenario 3). Follow-up data from the truncated troglitazone arm of the Diabetes Prevention Program (DPP) suggested that scenario 2 may be the case for thiazolidinediones (5).

We conducted a prospective 2- to 3-month post-trial medication washout to evaluate the impact on the new-onset type 2 diabetes rate in consenting DREAM trial participants who at trial end had not developed diabetes, were taking their double-blind study medication, and underwent an oral glucose tolerance test (OGTT).

RESEARCH DESIGN AND

METHODS—The design and primary results of the DREAM trial have been published (1-3). Between July 2001 and August 2003, 5,269 participants aged 30 years or more with impaired fasting plasma glucose (FPG; ≥6.1 but <7.0 mmol/L) and/or IGT (2-h post 75 g oral glucose load plasma glucose ≥7.8 but <11.1 mmol/L) were allocated at random to receive ramipril (titrated to a maximum of 15 mg) or matching placebo and, simultaneously, rosiglitazone (titrated to a maximum of 8 mg) or matching placebo, in a two-by-two balanced factorial design. Participants had OGTTs done after 2 years and at final visit, and at other yearly visits if FPG or HbA_{1c} values were elevated (2), and were followed for a median of 3 years. All participants provided informed written consent for the whole study, including the washout phase.

Consenting participants who underwent an OGTT at their last study visit, who had not developed diabetes, and who were taking their study medication were entered into the post-trial washout. They were given single-blind placebo rosiglitazone and ramipril medication and scheduled for an OGTT 2 to 3 months later. No other clinical assessments were done at the end of this period. The primary composite outcome for

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DREAM trial drug washout

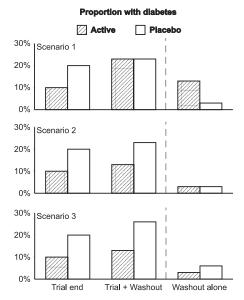


Figure 1—Three possible scenarios showing the relative proportions of subjects with new-onset diabetes at the end of a diabetes prevention trial, after the trial plus a drug washout period, and after the washout period alone. Scenario 1 illustrates a treatment that merely masks the appearance of diabetes by suppressing glucose levels. Scenario 2 demonstrates slowing the development of diabetes while the treatment is being given. Scenario 3 shows sustained benefit even after the treatment has been withdrawn.

the washout was the same as the main trial (new-onset diabetes or death), but the diagnosis of diabetes was based on one. not two, successive abnormal OGTTs or FPG values ≥7.0 mmol/L. As before, development of diabetes in an individual

was accepted if a physician outside the study diagnosed diabetes on the basis of a FPG level ≥7.0 mmol/L or a non-FPG level ≥11.1 mmol/L and had prescribed an antidiabetic agent.

Statistical analysis

Data were collected and analyzed at the Population Health Research Institute, McMaster University, using an intentionto-treat approach according to marginal groups: ramipril versus placebo and rosiglitazone versus placebo. Participants who entered the washout phase but who did not return for an OGTT and whose postwashout diabetes status was unknown were assumed not to have developed diabetes. Differences in trial-end prewashout characteristics by prior randomization were compared using t tests for continuous variables and χ^2 tests for categorical variables. Hazard ratios (HRs) and 95% CIs for the effect of prior treatment with each study drug (stratified by the other drug) on the primary outcome, and on the secondary outcome of regression to normoglycemia, were calculated using Cox proportional hazards models. Possible statistical interactions between ramipril and rosiglitazone treatments were assessed by the inclusion of an interaction term in the models. Trial plus washout analyses were done for all randomized participants from the time of randomization until the end of the washout period. Participants who did not enter the washout phase were analyzed according to their trial-end glycemic status. Additional washout alone analyses were restricted to those participants who entered the washout period. Median changes in fasting and 2-h plasma glucose levels during the washout phase were compared according to prior treatment with each study drug or respective placebo using Wilcoxon signed rank tests (with no adjustment for multiple testing).

RESULTS

Subject disposition

Of the 5,269 randomized participants the primary outcome occurred by trial end in 995, comprising the 992 reported previously (2,3) plus 3 more with delayed reports. An additional 907 people were ineligible for the washout because they had no final-visit OGTT (n = 64) or were no longer taking study medication (n = 843). Of the eligible 3,367, 105 declined to participate. The 3,262 (96.9%) participants completing the washout period were followed for a median (interquartile range [IOR]) of 71 (63–86) days. When compared with the 105 individuals who declined participation, these individuals had a similar duration of trial follow-up, age, sex, weight, waist-to-hip ratio, blood pressure, and mean FPG level (P all > 0.1), but lower 2-h postchallenge glucose value (7.13 mmol/L vs. 7.77 mmol/L; P = 0.0024).

Trial-end characteristics

Table 1 lists the participant characteristics at trial end according to their prior allocation to ramipril or placebo and to rosiglitazone or placebo. These reflect the

Table 1—Demographic, clinical, and biochemical characteristics of the 3,367 subjects without diabetes who entered the drug washout period

| | All | Ramipril | Placebo | P | Rosiglitazone | Placebo | P |
|--------------------------------|------------------|------------------|------------------|----------|------------------|------------------|----------|
| n | | - | | | | | |
| Entering washout | 3,367 | 1,677 | 1,690 | | 1,833 | 1,534 | |
| Completing washout | 3,262 (96.9%) | 1,632 (97.3%) | 1,630 (96.5%) | | 1,773 (96.7%) | 1,489 (97.1%) | |
| Female | 1,897 (58.2%) | 972 (59.6%) | 925 (56.8%) | 0.11 | 1,003 (56.67%) | 894 (60.0%) | 0.046 |
| Follow-up to trial end (years) | 3.2 ± 0.5 | 3.2 ± 0.5 | 3.2 ± 0.5 | 0.28 | 3.2 ± 0.5 | 3.2 ± 0.5 | 0.89 |
| Mean age at trial end (years) | 58.1 ± 10.7 | 58.1 ± 10.8 | 58.0 ± 10.6 | 0.83 | 57.9 ± 10.5 | 58.3 ± 10.9 | 0.30 |
| FPG (mmol/L) | 5.56 ± 0.72 | 5.54 ± 0.70 | 5.59 ± 0.74 | 0.029 | 5.46 ± 0.69 | 5.69 ± 0.74 | < 0.0001 |
| 2HPG (mmol/L) | 7.13 ± 1.96 | 7.04 ± 1.94 | 7.22 ± 1.98 | 0.013 | 6.76 ± 1.82 | 7.57 ± 2.03 | < 0.0001 |
| BP (mmHg) | | | | | | | |
| Systolic | 129.2 ± 17.1 | 127.0 ± 17.1 | 131.5 ± 16.8 | < 0.0001 | 128.4 ± 16.8 | 130.2 ± 17.4 | 0.0025 |
| Diastolic | 78.4 ± 10.6 | 77.2 ± 10.7 | 79.6 ± 10.3 | < 0.0001 | 77.7 ± 10.4 | 79.2 ± 10.6 | < 0.0001 |
| BMI (kg/m^2) | 31.0 ± 5.7 | 31.0 ± 5.7 | 31.0 ± 5.7 | 0.94 | 31.6 ± 5.7 | 30.4 ± 5.6 | < 0.0001 |
| Weight (kg) | 85.3 ± 19.4 | 85.1 ± 19.3 | 85.5 ± 19.5 | 0.56 | 87.3 ± 19.7 | 83.0 ± 18.7 | < 0.0001 |
| Waist/hip | | | | | | | |
| Male | 0.96 ± 0.08 | 0.96 ± 0.08 | 0.96 ± 0.08 | 0.58 | 0.96 ± 0.08 | 0.96 ± 0.07 | 0.62 |
| Female | 0.87 ± 0.12 | 0.87 ± 0.13 | 0.87 ± 0.10 | 0.60 | 0.86 ± 0.13 | 0.88 ± 0.11 | 0.0075 |

Categorical variables are expressed as n (%) and continuous variables as mean \pm 1 SD.

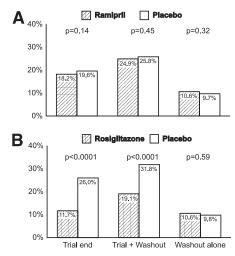


Figure 2—Proportions of subjects with newonset diabetes at the end of the trial, after the trial plus drug washout, and after the washout period (median 71 days) alone for ramipril vs. placebo (A) and rosiglitazone vs. placebo arms of the trial (B).

glycemic impact of these therapies before commencing the washout phase. Participants allocated previously to ramipril (n =1,632), compared with those allocated to placebo (n = 1,630), had lower fasting and 2-h postchallenge plasma glucose values and lower systolic and diastolic blood pressure values. Participants allocated previously to rosiglitazone (n =1,773), compared with those allocated to placebo (n = 1.489), had lower fasting and 2-h postchallenge plasma glucose values and lower systolic and diastolic blood pressure values. They were also less likely to be women and were heavier, and the females had a lower waist-to-hip ratio.

Ramipril outcomes

At trial-end plus washout, the proportion of participants in whom the primary outcome had occurred (Fig. 2A and Table 2) did not differ between those allocated to

ramipril or to placebo (24.9 vs. 25.8%, respectively; P = 0.45); the proportion of participants regressing to normoglycemia was greater (HR 1.11 [95% CI 1.01–1.21]; P = 0.031) for those allocated to ramipril (36.0%) compared with placebo (33.2%). During the washout period alone, both the primary and secondary outcome occurred at the same rate in both groups (10.6 vs. 9.7%, P = 0.32; and 23.0 vs. 22.7%, P = 0.66, respectively).

Rosiglitazone outcomes

At the end of the trial plus washout, the proportion of participants in whom the primary outcome had occurred (Fig. 2B and Table 2) was less (HR 0.51 [95% CI 0.45-0.57]; P < 0.0001) for those allocated to rosiglitazone (19.0%) compared with placebo (31.7%); the proportion of participants regressing to normoglycemia was greater (HR 1.22 [1.11–1.34]; P <0.0001) for those allocated to rosiglitazone (39.2%) compared with placebo (30.0%). During the washout period alone, both the primary and secondary outcome occurred at the same rate in both groups (10.5 vs. 9.8%, P = 0.59; and 21.5 vs. 23.8%, P = 0.33, respectively).

Postwashout fasting and 2-h plasma glucose levels

At the end of the washout, there was no difference in either fasting or 2-h post-challenge median plasma glucose levels between people originally allocated to ramipril versus placebo (Table 3). At the same time median FPG levels did not differ between those allocated to rosiglitazone or placebo, despite lower FPG levels in the rosiglitazone group at trial end (Table 3). This was the result of a greater increase in FPG levels during the washout in participants formerly on rosiglitazone versus placebo (0.30 vs. 0.10 mmol/L, respectively; P < 0.0001).

Conversely, 2-h postchallenge plasma glucose levels remained lower after the washout (7.1 vs. 7.4 mmol/L; P = 0.005) despite a greater increase in the rosiglitazone group (0.5 mmol/L) compared with the placebo group (0.0 mmol/L; P < 0.0001).

Interaction effects

No statistical interactions were observed between ramipril and rosiglitazone with respect to FPG levels, 2-h plasma glucose levels, or glycemic status at the end of the trial plus washout or during the washout period alone.

Further analyses

No differences were seen in the results for individuals entering the study with IFG as opposed to IGT.

CONCLUSIONS—The DREAM study primary outcome (new-onset diabetes or death) results remained essentially similar when reanalyzed after a median 71-day study medication washout period. With respect to the washout period itself, because ramipril did not reduce diabetes incidence in the main trial (2) it is not surprising that the postwashout newonset diabetes rates did not differ between those allocated ramipril or placebo during the trial. The lack of any differential impact on post-trial diabetes incidence between those allocated rosiglitazone or placebo during the trial, despite the major within-trial reduction in diabetes incidence (3), suggests that this agent does not have a sustained effect on the underlying disease pathophysiology.

The incidence of diabetes after therapy cessation with either of the agents evaluated did not differ from the incidence observed with placebo. These findings suggest that the underlying disease process is slowed only while the intervention is being given (scenario 2). The

Table 2—HRs (95% CI) for development of diabetes or death at trial end, at trial end plus washout, and during washout period alone for the ramipril vs. placebo and for the rosiglitazone vs. placebo arms of the trial

| | Ramipril | Placebo | HR (95% CI) | P | Rosiglitazone | Placebo | HR (95% CI) | P |
|--------------------|-------------|-------------|------------------|--------|---------------|-----------|------------------|----------|
| Primary outcome | | | | | | | | |
| Trial end | 476/2,623 | 519/2,646 | 0.91 (0.80-1.03) | 0.14 | 309/2,635 | 686/2,634 | 0.40 (0.35-0.46) | < 0.0001 |
| Trial + washout | 654/2,623 | 683/2,646 | 0.95 (0.84-1.08) | 0.45 | 501/2,635 | 836/2,634 | 0.51 (0.45-0.57) | < 0.0001 |
| Washout alone | 178/1,677 | 164/1,690 | 1.12 (0.90-1.38) | 0.32 | 192/1,833 | 150/1,534 | 1.06 (0.86-1.32) | 0.59 |
| Secondary outcome* | | | | | | | | |
| Trial end | 1,117/2,623 | 1,013/2,646 | 1.17 (1.07-1.27) | 0.0002 | 1,329/2,635 | 801/2,634 | 1.69 (1.55-1.85) | < 0.0001 |
| Trial + washout | 944/2,623 | 878/2,646 | 1.11 (1.01-1.21) | 0.031 | 1,032/2,635 | 790/2,634 | 1.22 (1.11-1.34) | < 0.0001 |
| Washout alone | 164/714 | 183/807 | 1.05 (0.85–1.30) | 0.66 | 144/669 | 203/852 | 0.90 (0.73–1.12) | 0.33 |

^{*}Regression to normoglycemia.

DREAM trial drug washout

Table 3—Median (IQR) fasting and 2-h plasma glucose levels (mmol/L) at trial end, at trial end plus washout, and during washout period alone for the ramipril vs. placebo and for the rosiglitazone vs. placebo arms of the trial

| | Ramipril | | | Placebo | | | | R | Rosiglitazone | | | Placebo | | |
|-----------------|----------|--------|---------|---------|--------|---------|-------|-------|---------------|---------|-------|---------|---------|----------|
| | n | Median | IQR | n | Median | IQR | P | n | Median | IQR | n | Median | IQR | P |
| Fasting glucose | | | | | | | | | | | | | | |
| Baseline | 2,623 | 5.9 | 5.4-6.3 | 2,646 | 5.9 | 5.4-6.4 | 0.45 | 2,635 | 5.9 | 5.4-6.4 | 2,634 | 5.9 | 5.4-6.3 | 0.52 |
| Trial end | 2,418 | 5.7 | 5.2-6.3 | 2,440 | 5.7 | 5.2-6.5 | 0.033 | 2,431 | 5.5 | 5.1-6.1 | 2,427 | 5.9 | 5.3-6.5 | < 0.0001 |
| Washout | 1,630 | 5.7 | 5.2-6.2 | 1,628 | 5.7 | 5.3-6.2 | 0.090 | 1,769 | 5.7 | 5.2-6.2 | 1,489 | 5.7 | 5.2-6.2 | 0.72 |
| 2-h Glucose | | | | | | | | | | | | | | |
| Baseline | 2,623 | 8.6 | 8.0-9.6 | 2,646 | 8.8 | 8.0-9.8 | 0.064 | 2,635 | 8.7 | 8.0-9.7 | 2,634 | 8.7 | 8.0-9.7 | 0.85 |
| Trial end | 2,032 | 7.1 | 5.8-8.7 | 2,033 | 7.3 | 5.9-8.9 | 0.020 | 2,168 | 6.7 | 5.6-8.3 | 1,897 | 7.7 | 6.2-9.4 | < 0.0001 |
| Washout | 1,624 | 7.2 | 5.9-8.6 | 1,610 | 7.3 | 6.0-8.9 | 0.052 | 1,754 | 7.1 | 5.9-8.6 | 1,480 | 7.4 | 6.0-9.0 | 0.0051 |

fact that most (but not all) of their effects on glucose levels disappear with drug discontinuation also suggests that underlying β -cell defects are not reversed or repaired and that the minor metabolic effects of ramipril and the profound effects of rosiglitazone are likely secondary to increased sensitivity to insulin.

The rosiglitazone results are consistent with the observations made in a similar population of 585 people with IGT who took troglitazone for a median period of 0.9 years in the DPP trial (5). When compared with placebo, troglitazone reduced the incidence of diabetes by 75% while it was being taken, but when discontinued (because of emerging concerns about liver toxicity) diabetes incidence was similar to, but not greater than, that in the placebo group. Therefore, in the context of the full 3-year follow of the DPP, study participants who had taken troglitazone for 0.9 years were only $\sim 17\%$ less likely to have developed diabetes. By contrast, these findings are inconsistent with a study in women with a history of gestational diabetes (4) where troglitazone reduced the incidence of diabetes by 55% compared with placebo while it was being taken, but the effect persisted during an 8-month washout period. Postwashout glucose tolerance testing, however, was only done in a subset of randomized people who completed the trial without diabetes and who then returned after the washout. The possibility that responders were more likely to return for reassessment than nonresponders cannot be excluded.

The prospectively planned DREAM washout was completed while participants were taking single-blind placebo medication. Moreover, 97% of eligible participants completed this phase of the study with completion rates the same

in those previously on active therapies or placebo (Table 1). These strengths support the robustness of the DREAM washout findings but there are several limitations. First, the ~10% primary outcome incidence during a median washout of 71 days is higher than expected given the 26% 3-year incidence in the rosiglitazone placebo group during the trial (Fig. 2B). This is likely because of the fact that during the washout phase incident diabetes was based on only a single abnormal fasting or 2-h plasma glucose level as opposed to two successive abnormal values during the trial. Because of regression to the mean (7), many of those classified with diabetes based on a single abnormal glucose value during washout would not have had this diagnosis confirmed on a second test. Similar discrepancies in post-trial diabetes incidence rates have been noted in other diabetes prevention trials, which reported high diabetes incidence rates based on one abnormal glucose value following a short washout phase (5,8). Any overestimate of the absolute incidence of diabetes would, however, affect all treatment groups to the same extent so should not invalidate between-group comparisons. A second limitation is the relatively short washout period of a median 71 days. This may have been insufficient to totally wash out rosiglitazone, which may require a longer period of study medication. Finally, the effect of the washout on other secondary measurements such as liver function tests, blood pressure, body weight, and edema was not assessed.

In summary, rosiglitazone delays disease progression during treatment, but the process resumes at the placebo rate when the drug is stopped, as has been seen previously with metformin (9). Additional longer-term assessments of glucose

tolerance in epidemiologic follow-up studies may yield further insights in the context of the effect of these interventions on the natural history of IFG, IGT, and type 2 diabetes. At this time the data indicate that rosiglitazone can substantially reduce the incidence of type 2 diabetes while it is being administered, but this effect is not sustained when the drug is withdrawn.

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R.R.H. wrote the manuscript and researched data. B.Z. contributed to discussion and reviewed and edited the manuscript. S.Y. reviewed and edited the manuscript. P.M.S. and S.S.A. contributed to discussion. J.J.B. researched data and reviewed and edited the manuscript. I.C., M.J.D., and V.P. reviewed and edited the manuscript. G.R.D. researched data and reviewed the manuscript. J.M.P. researched data and reviewed and edited the manuscript. P.Z.Z. reviewed and edited the manuscript. H.C.G. researched data and wrote the manuscript.

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