# Lipoprotein(a) Is an Independent Risk Factor for Peripheral Arterial Disease in Chinese Type 2 Diabetic Patients in Taiwan 

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#### Abstract

OBJECTIVE - To examine the association between lipoprotein(a) [Lp(a)] and peripheral arterial disease (PAD) and determine the optimal cutoff in Chinese type 2 diabetic patients in Taiwan.


#### Abstract

RESEARCH DESIGN AND METHODS - Serum Lp(a) was determined in 557 type 2 diabetic patients ( 243 men and 314 women) recruited consecutively from a diabetes clinic at the National Taiwan University Hospital. Ankle-brachial index (ABI) <0.9 was diagnosed as PAD $(n=45)$ and $<0.8$ as severe PAD $(n=20)$. Potential confounders included age, sex, BMI, waist-to-hip ratio (WHR), diabetes duration, insulin usage, smoking, hypertension, systolic and diastolic blood pressure, fasting plasma glucose (FPG), total cholesterol, triglycerides, and HDL and LDL cholesterol.

RESULTS - The distribution of $\operatorname{Lp}(a)$ was right skewed and no significant differences for sex, WHR, insulin usage, smoking, hypertension, and systolic and diastolic blood pressure were observed. In men, $\log [\operatorname{Lp}(a)]$ was correlated positively with age, duration, and total and LDL cholesterol (borderline significant, $P<0.1$ ) and negatively with BMI, triglycerides, and FPG ( $P<0.1$ ). In women, $\log [\operatorname{Lp}(\mathrm{a})]$ was correlated positively with total and LDL cholesterol and negatively with triglycerides and BMI ( $P<0.1$ ). ABI was significantly correlated with $\log [\operatorname{Lp}(a)]$, especially in men or in patients with PAD. The optimal cutoff determined by discriminant analysis was $13.3 \mathrm{mg} / \mathrm{dl}$. Patients with $\operatorname{Lp}(\mathrm{a})$ above this value had a 2.7-fold higher risk of PAD after multivariate adjustment. $\operatorname{Lp}(a)$ also significantly increased from no PAD to mild and severe $\operatorname{PAD}(17.1 \pm 14.4,23.7 \pm 20.3$, and $36.9 \pm 22.8 \mathrm{mg} / \mathrm{dl}$, respectively, $P<0.001)$.


CONCLUSIONS - Lp (a) is an independent risk factor for PAD in type 2 diabetic patients in Taiwan. The optimal cutoff is $13.3 \mathrm{mg} / \mathrm{dl}$.

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Lipoprotein(a) [Lp(a)] is a heterogeneous macromolecule consisting of a glycoprotein apolipoprotein(a), which is disulfide-linked to apolipoprotein B-100 on an LDL core (1). Apoli-
poprotein(a) exhibits size polymorphism, which is closely linked to $\operatorname{Lp}(a)$ density and concentrations (2). The limited distribution of $\operatorname{Lp}(a)$ in a few animal species implies that it is not essential in lipopro-

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Abbreviations: ABI, ankle-brachial index; FPG, fasting plasma glucose; Lp(a), lipoprotein(a); PAD, peripheral artery disease; WHR, waist-to-hip ratio.

A table elsewhere in this issue shows conventional and Systeme International (SI) units and conversion factors for many substances.
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tein metabolism (3). However, it is clinically important because its concentrations are primarily genetically determined, associated with atherosclerotic disease, and less affected by lifestyle or medication (4). Lp(a) concentrations are quite constant in an individual (2). Lipid-lowering statins appear ineffective, whereas niacin might have some effect on $\operatorname{Lp}(a)$ lowering (5). Lp(a) might be higher in type 2 diabetic patients, but glycemic control seems to have no effect on serum $\operatorname{Lp}(a)$ (6). With few exceptions $(7,8), \operatorname{Lp}(a)$ has been shown to be a risk factor for atherosclerotic disease, such as ischemic heart disease $(9,10)$, myocardial infarction $(11,12)$, stroke (13), and peripheral arterial disease (PAD) $(14,15)$.

A community study in Taiwan examining the relationship between $\operatorname{Lp}(a)$ and socioeconomic and atherosclerotic risk factors demonstrated that $\mathrm{Lp}(\mathrm{a})$ was positively correlated with age and LDL and HDL cholesterol and negatively correlated with triglycerides, obesity, and insulin resistance (16). Socioeconomic status, smoking, alcohol consumption, systolic and diastolic blood pressure, and apolipoprotein Al and B did not correlate with $\operatorname{Lp}(a)$ (16). Although this study demonstrated the association between $\operatorname{Lp}(a)$ and some of the atherosclerotic risk factors, whether $\operatorname{Lp}(a)$ is an independent risk factor for atherosclerotic disease in the Chinese population in Taiwan awaits further evaluation. Most studies evaluating Lp(a) and atherosclerotic disease focused on coronary artery disease, and only a few evaluated its association with PAD. Data in the Asian populations are especially rare, and nothing is known about Lp(a) and PAD in the Chinese population in Taiwan. It is not known whether the recommended cutoff of $30.0 \mathrm{mg} / \mathrm{dl}$ for Caucasians is appropriate for Chinese type 2 diabetic patients in Taiwan. Thus, the objectives of this study were to 1) examine whether $\operatorname{Lp}(a)$ was an independent risk factor for PAD in Chinese type 2 diabetic
patients and 2) determine the optimal cutoff of $\mathrm{Lp}(\mathrm{a})$ for discriminating patients with and without PAD.

## RESEARCH DESIGN AND

METHODS - A total of 557 type 2 diabetic patients ( 243 men and 314 women) were consecutively recruited from a diabetes clinic at the National Taiwan University Hospital in Taipei, Taiwan. The patients were treated with either oral antidiabetic drugs or insulin at recruitment. They did not show a history of diabetic ketoacidosis at the onset of diabetes or receive insulin treatment within 1 year of diagnosis. Patients with acute illness or taking niacin, estrogen replacements, or antibiotics were not included.

## Diagnosis of PAD

Diagnosis of PAD was based on an anklebrachial index (ABI) $<0.9$ on either leg as described in previous studies (17-19). In brief, Doppler ultrasound (Medacord PVL; MedaSonic, Mountain View, CA) was used to measure the systolic pressures on bilateral brachial, posterior tibial, and dorsal pedal arteries. The device calculated the right and left ABI automatically by dividing the higher pressure on the dorsal pedal or posterior tibial arteries on right and left sides, respectively, by the higher brachial pressure on either side. PAD was further categorized as a severe form (ABI $<0.80$ on either side of the extremities) or as a mild form (ABI $<0.90$ on either leg, but neither $<0.80$ ).

## Risk factors

Risk factors included age, sex, BMI, waist-to-hip ratio (WHR), diabetes duration, hypertension, smoking, insulin usage, systolic blood pressure, diastolic blood pressure, fasting plasma glucose (FPG), serum total cholesterol, triglycerides, HDL cholesterol, LDL cholesterol, and Lp(a). While evaluating the association between $\operatorname{Lp}(a)$ and $P A D$, the other risk factors were treated as potential confounders. Blood pressure was measured on the right arm after a $20-\mathrm{min}$ rest in a sitting position with a mercury sphygmomanometer by the auscultatory method. The first perception of successive sounds (Korotkoff phase I) was taken as systolic blood pressure, and the complete disappearance of sound (Korotkoff phase V) was taken as diastolic blood pressure. Hypertension was defined as systolic blood pressure $\geq 140 \mathrm{mmHg}$ and/or diastolic
blood pressure $\geq 90 \mathrm{mmHg}$ (20) or if the patient was receiving antihypertensive therapy. Venous blood samples were collected in the morning after an overnight fast of $>12 \mathrm{~h}$. Serum samples were used to determine total cholesterol, triglycerides, and HDL and LDL cholesterol on an autoanalyzer (Hitachi 737) with reagents obtained from Boehringer Mannheim Diagnostics (Indianapolis, IN) $(17,18)$. Serum $\operatorname{Lp}(a)$ was measured by turbidimetric immunoassay method [Biolatex Lp(a) assay; Biolatex, Logroño, Spain]. A 10-fold dilution was made before assay if the serum sample was turbid or if the triglyceride level was $>400 \mathrm{mg} / \mathrm{dl}$. FPG was determined by a glucose oxidase method (21). Measurements of anthropometric parameters are described elsewhere (22).

## Statistical analyses

Because the distribution of $\operatorname{Lp}(a)$ was highly right skewed, the natural logarithm $\{\log [\operatorname{Lp}(a)]\}$ was used for statistical tests when an assumption of normal distribution was required. The differences of $\log [\operatorname{Lp}(a)]$ among different age-groups were compared by one-way ANOVA in separate sexes. Student's $t$ test was used to test the differences of $\log [\operatorname{Lp}(a)]$ between the diabetic men and women in each stratum of age and in all ages, between patients with and without hypertension, between insulin users and nonusers, and between smokers and nonsmokers. Pearson correlation coefficients were generated to evaluate the linear relationship between continuous covariates and lo$g[\operatorname{Lp}(a)]$ in separate sexes.
$\log [\operatorname{Lp}(a)]$ was used in discriminant analysis (Wilks' method) to determine the optimal cutoff of Lp(a). Patients were divided into different subgroups by using tertiles, quartiles, and quintiles of $\operatorname{Lp}(a)$, by using the cutoff of $30.0 \mathrm{mg} / \mathrm{dl}$, and by the cutoff derived from the discriminant analysis. Prevalences of PAD in different subgroups of $\operatorname{Lp}(a)$ were analyzed by $\chi^{2}$ test. Logistic regression was used to estimate the multivariate-adjusted odds ratios (ORs) and their 95\% CIs for PAD among the different subgroups by using the lowest $\operatorname{Lp}(a)$ subgroups as reference groups. All potential confounders, including age, sex, BMI, WHR, diabetes duration, hypertension, smoking, insulin usage, systolic blood pressure, diastolic blood pressure, FPG, total cholesterol, triglycerides, and HDL and LDL cholesterol, were adjusted for in these models.

Table 1—Baseline clinical characteristics of the subjects

| $n$ | 557 |
| :--- | :---: |
| Age (years) | $63.4 \pm 10.4$ |
| Sex (\% men) | 43.6 |
| BMI $\left(\mathrm{kg} / \mathrm{m}^{2}\right)$ | $24.7 \pm 3.5$ |
| WHR | $0.90 \pm 0.08$ |
| Diabetes duration (years) | $12.1 \pm 7.9$ |
| Insulin user | 20.1 |
| Smoker | 27.8 |
| Hypertension | 37.9 |
| Systolic blood pressure (mmHg) | $134.7 \pm 17.5$ |
| Diastolic blood pressure (mmHg) | $79.0 \pm 9.7$ |
| FPG (mg/dl) | $152.0 \pm 51.3$ |
| Total cholesterol (mg/dl) | $202.3 \pm 38.8$ |
| Triglyceride (mg/dl) | $174.1 \pm 112.8$ |
| HDL cholesterol (mg/dl) | $49.7 \pm 33.9$ |
| LDL cholesterol (mg/dl) | $112.6 \pm 48.5$ |
| Lp(a) (mg/dl) | $18.1 \pm 15.5$ |
| Data are means $\pm$ SD or percent. |  |

Lp(a) was also treated as a continuous variable while estimating the OR by logistic regression. Various model-based ORs were created, either before or after adjustment for confounders. Because age and sex are unmodifiable risk factors, respective models adjusted for age, sex, and both age and sex were first created to see the possible influence of these two factors on the estimated ORs. Further models were created by selecting, in addition to age and sex, 1) a variable with $P<0.1$ in any univariate analysis (i.e., age and sex plus one of the following variables at a time: BMI, diabetes duration, FPG, total cholesterol, triglycerides, or LDL cholesterol), 2) all variables with $P<0.1$ in any univariate analysis (i.e., age and sex plus all of the following variables: BMI, diabetes duration, FPG, total cholesterol, triglycerides, and LDL cholesterol), and 3) all potential confounders (i.e., age, sex, BMI, WHR, diabetes duration, hypertension, smoking, insulin usage, systolic blood pressure, diastolic blood pressure, FPG, total cholesterol, triglycerides, HDL cholesterol, and LDL cholesterol). All continuous variables were not categorized while used for adjustment. One-way ANOVA was used to compare the differences of $\log [\operatorname{Lp}(a)]$ among the three subgroups of patients in respect to PAD , i.e., patients without, with mild, and with severe PAD.

Whenever there was a significant $P$ value of $<0.05$ in one-way ANOVA, multiple comparisons by least significant difference were further performed. Continuous variables were expressed as
means $\pm$ SD. $P<0.05$ was considered statistically significant, and $0.05<P<$ 0.1 was considered borderline significant.

RESULTS - The baseline characteristics of the subjects are shown in Table 1. $\log [\operatorname{Lp}(a)]$ values were not significantly different among the different age-groups ( $<45,45-54,55-64,65-74$, and $\geq 75$ years) in either sex and between sexes in each stratum of age. Lp(a) levels in the diabetic men and women were $18.2 \pm$ 16.2 and $18.1 \pm 15.0 \mathrm{mg} / \mathrm{dl}$, respectively $\{$ Student's $t$ test for $\log [\operatorname{Lp}(a)], P>0.1\}$. The differences of $\log [\operatorname{Lp}(a)]$ were not statistically significant between patients with and without hypertension, smokers and nonsmokers, and insulin users and nonusers (data not shown). However, patients with PAD had a significantly higher level of $\operatorname{Lp}(a)$ than patients without PAD $\{29.5 \pm 22.2$ vs. $17.1 \pm 14.4 \mathrm{mg} / \mathrm{dl}$, Student's $t$ test for $\log [\operatorname{Lp}(a)], P<0.001\}$.

Table 2 shows the Pearson correlation coefficients between $\log [\operatorname{Lp}(a)]$ and continuous covariates in separate sexes and in total patients and between $\log [\operatorname{Lp}(a)]$ and ABI in all patients, patients with PAD , and patients without PAD, respectively.

The cutoff of $\operatorname{Lp}(a)$ generated from the discriminant analysis was $13.3 \mathrm{mg} / \mathrm{dl}$, which was approximately equal to the median value of $13.1 \mathrm{mg} / \mathrm{dl}$. Sixty percent of the originally grouped cases with and without PAD would be correctly classified by this cutoff. Prevalences of and multi-

Table 3—Prevalence and multivariate-adjusted ORs for PAD in different cutoffs of Lp(a)

|  | Lp(a) (mg/dL) | PAD prevalence | Adjusted OR (95\% CI) |
| :--- | :---: | :---: | :---: |
| Cutoff points (mg/dl) |  |  |  |
| $\quad<30$ | $11.68 \pm 7.51$ | 5.8 | 1.000 |
| $\geq 30$ | $44.40 \pm 11.50$ | $17.3^{*}$ | $3.019(1.433-6.360)^{*}$ |
| $<13.3$ | $6.85 \pm 3.43$ | 4.9 | 1.000 |
| $\geq 13.3$ | $29.96 \pm 14.42$ | $11.4^{*}$ | $2.685(1.203-5.994)^{*} \dagger$ |
| Tertiles |  |  |  |
| I | $4.84 \pm 2.27$ | 5.4 | 1.000 |
| II | $13.33 \pm 3.18$ | 5.4 | $1.111(0.418-2.951)$ |
| III | $36.22 \pm 13.37$ | $13.4^{*}$ | $2.460(1.044-5.798)^{*} \dagger$ |
| Quartiles |  |  |  |
| I | $3.89 \pm 1.81$ | 3.6 | 1.000 |
| II | $9.49 \pm 1.74$ | 6.5 | $1.558(0.465-5.216)$ |
| III | $18.29 \pm 3.50$ | 6.4 | $1.740(0.519-5.832)$ |
| IV | $40.88 \pm 12.33$ | $15.8^{*}$ | $4.666(1.531-14.221)^{*}$ |
| Quintiles |  |  |  |
| I | $3.28 \pm 1.50$ | 2.7 | 1.000 |
| II | $7.82 \pm 1.19$ | 7.1 | $1.967(0.460-8.413)$ |
| III | $13.09 \pm 2.12$ | 5.4 | $2.085(0.461-9.423)$ |
| IV | $22.29 \pm 3.79$ | 7.1 | $2.752(0.642-11.797)$ |
| V | $44.26 \pm 11.53$ | $18.0^{*}$ | $6.333(1.629-24.612)^{*}$ |
| Continuous |  |  | $1.043(1.021-1.066)^{*}$ |

Data are means $\pm$ SD or percent. $* P<0.01 ; \dagger P<0.05$.
variate-adjusted OR for PAD in subgroups of tertiles, quartiles, quintiles, and cutoffs of $30.0 \mathrm{mg} / \mathrm{dl}$ and $13.3 \mathrm{mg} / \mathrm{dl}$ of Lp(a) are shown in Table 3. The OR for PAD after adjustment for all potential confounders when $\operatorname{Lp}(a)$ was treated as a continuous variable is also given in Table 3. While $\operatorname{Lp}(a)$ was treated as a continu-

Table 2—Pearson correlation coefficients between $\log [\operatorname{Lp}(a)]$ and continuous covariates and ABI by sex

| Variables | Men | Women | Total |
| :--- | :---: | :---: | :---: |
| $n$ | 243 | 314 | 557 |
| Age | $0.150^{*}$ | 0.040 | $0.092^{*}$ |
| BMI | $-0.164^{*}$ | $-0.108 \dagger$ | $-0.133 \ddagger$ |
| WHR | 0.047 | -0.059 | -0.023 |
| Diabetes duration | $0.131^{*}$ | 0.012 | 0.065 |
| Systolic blood pressure | 0.011 | -0.001 | 0.006 |
| Diastolic blood pressure | -0.084 | -0.019 | -0.046 |
| FPG | $-0.121 \dagger$ | 0.026 | -0.041 |
| Total cholesterol | $0.156^{*}$ | $0.214 \ddagger$ | $0.189 \ddagger$ |
| Triglyceride | $-0.138^{*}$ | $-0.111^{*}$ | $-0.120 \ddagger$ |
| HDL cholesterol | 0.077 | 0.023 | 0.034 |
| LDL cholesterol | $0.125^{*} \dagger$ | $0.199 \ddagger$ | $0.169 \ddagger$ |
| ABI |  |  |  |
| $\quad$ All patients | $-0.261 \ddagger$ | $-0.108 \dagger$ | $-0.185 \ddagger$ |
| Patients with PAD | $-0.108 \dagger(n=16)$ | $-0.176(n=29)$ | $-0.340^{*}(n=45)$ |
| Patients without PAD | $-0.101(n=227)$ | $-0.019(n=285)$ | $-0.073(n=512)$ |

*P $<0.05 ; \dagger 0.05<P<0.1 ; \ddagger P<0.01$.
ous variable, the various model-based ORs for PAD were all highly significant ( $P<0.001$ ) with similar values either before or after various adjustments (data not shown). The percentile distributions of $\operatorname{Lp}(a)$ in patients without, with mild ( $n=$ $25)$, and with severe $(n=20)$ PAD are shown in Fig. 1. Mean values of $\operatorname{Lp}(a)$ in these subgroups of patients are shown in Fig. 2.

CONCLUSIONS - This study confirmed the association between $\operatorname{Lp}(a)$ and PAD, independent of potential confounders, in Chinese type 2 diabetic patients in Taiwan. The conventionally used cutoff of $30.0 \mathrm{mg} / \mathrm{dl}$, which corresponded to approximately the 80th percentile of this sample, carried a three times higher risk of PAD (Table 3). However, the optimal cutoff in Chinese type 2 diabetic patients in Taiwan seemed to be at a lower level $(13.3 \mathrm{mg} / \mathrm{dl})$, which gave the greatest correct predictions ( $60 \%$ ) for patients with and without PAD and significantly increased the risk of PAD by 2.7-fold (Table 3). About $40 \%$ of the originally grouped cases with and without PAD would not be correctly classified by this cutoff, suggesting that some other risk factors would also be important to the development of PAD in type 2 diabetic patients. Thus, one


Figure 1—The percentile distributions of $\operatorname{Lp}(a)$ in patients without ( $\mathrm{n}=512$, $\bigcirc$ ), with mild ( $\mathrm{n}=$ 25 , $\boldsymbol{\square})$ and severe $(\mathrm{n}=20, \mathbf{\Delta})$ PAD.
should not neglect the possible existence of PAD in patients with a lower level of Lp(a).

In this study, $\operatorname{Lp}(a)$ was not only an independent risk factor for PAD, it was also correlated with ABI, especially when PAD was present (Table 2). Lp(a) also increased steadily from absence of PAD to mild and severe PAD (Fig. 2).

In the general population of Taiwan, age is a major determinant of $\operatorname{Lp}(a)$ in both sexes, and women tend to have slightly higher $\operatorname{Lp}(a)$ than their male counterparts (16). The proportion of people with $\mathrm{Lp}(\mathrm{a}) \geq 30.0 \mathrm{mg} / \mathrm{dl}$ is also higher in women in the general population (men versus women, 11.6 vs. $14.3 \%, P<0.05$ ) (16). However, these levels were not observed in the diabetic patients in this study. The correlation between Lp(a) and age was only observed in the diabetic men and not in the diabetic women (Table 2). The diabetic men and women have similar levels of $\operatorname{Lp}(a)$ (men versus women, $18.2 \pm 16.2$ vs. $18.1 \pm 15.0 \mathrm{mg} / \mathrm{dl}, P>$ $0.1)$ and similar proportions of $\operatorname{Lp}(a)$ $\geq 30.0 \mathrm{mg} / \mathrm{dl}$ (men versus women, 19.8
vs. $19.7 \%, P>0.1)$. These observations suggest that a higher proportion of diabetic patients would have $\operatorname{Lp}(a) \geq 30.0$ $\mathrm{mg} / \mathrm{dl}$ than the general population, without much influence by age or sex.

Similar to the observation in the general population (16), $\log [\operatorname{Lp}(a)]$ was negatively correlated with BMI and positively with total and LDL cholesterol and was not correlated with blood pressure (Table 2). In the diabetic patients, WHR and HDL cholesterol were not significantly correlated with $\log [\operatorname{Lp}(a)]$, and FPG was only negatively correlated with log [Lp(a)], with borderline significance in men (Table 2). The significant correlation between $\log [\operatorname{Lp}(a)]$ and triglycerides (negatively), as seen in the diabetic patients (Table 2), was also observed in the general population (16). Because Lp(a) is deemed a proatherogenic molecule, the negative association with BMI and triglycerides is puzzling and deserves further investigation. Similar to other studies (6), glycemic control seemed to have an insignificant effect on $\operatorname{Lp}(a)$ in type 2 diabetic patients in this study.


Figure 2—The mean values of $\operatorname{Lp}(a)$ in patients without, with mild, and with severe PAD (error bars indicate $S E)$. One-way ANOVA for $\log [\operatorname{Lp}(a)]$ among the three groups: $\mathrm{P}<0.001$. Multiple comparison test by least significant difference: no PAD vs. mild PAD, $\mathrm{P}<0.1$; no PAD vs. severe $P A D, \mathrm{P}<0.001$; mild $P A D$ vs. severe $P A D, 0.05<\mathrm{P}<0.1$.

Elevated total cholesterol and triglycerides are not good markers for PAD in the Chinese population in Taiwan, in either the diabetic patients (23) or subjects living in a confined community (18). This could partly be ascribed to the fact that although $\operatorname{Lp}(a)$ is less influenced by lifestyle and medication, both total cholesterol and triglycerides tend to be modified significantly by diet, exercise, and drug treatment. This study demonstrated that Lp(a) could be a much better marker for atherosclerotic disease in the Chinese population in Taiwan.

Although Lp(a) concentrations are primarily affected by genetic factors, recent studies (1) have demonstrated that its atherothrombogenic effect may be modified by oxidative events and actions of lipolytic and proteolytic enzymes. This may partly explain the reported discrepancy in the association between $\operatorname{Lp}(a)$ and cardiovascular disease in different studies. More research is needed to explore the molecular and pathogenic mechanisms related to the atherogenic effect of Lp(a).

## Limitations

This study was carried out in type 2 diabetic patients in a medical center in Taiwan. The possibility of referral bias cannot be ruled out. Furthermore, an association found in a cross-sectional study is not sufficient to conclude a causal relationship. It is also unknown whether $\mathrm{Lp}(\mathrm{a})$ is a risk factor for atherosclerotic diseases other than PAD. The usefulness of the optimal cutoff of $\operatorname{Lp}(a)$ found in this study also requires further investigation.

## Summary

$\mathrm{Lp}(\mathrm{a})$ is a significant and independent risk factor for PAD in type 2 diabetic patients in Taiwan. It is also significantly predictive for the severity of PAD, especially in patients with existent PAD or an ABI $<0.9$. The optimal cutoff of $\operatorname{Lp}(a)$ for PAD in type 2 diabetic patients in Taiwan is $13.3 \mathrm{mg} / \mathrm{dl}$, a level much lower than the commonly recommended $30.0 \mathrm{mg} / \mathrm{dl}$ in Caucasians.

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